

Multiple-output Composite Quantile Regression through an Optimal Transport Lens

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Problem setup

Parameter estimation problem

Linear model: $(X, Y) \in \mathbb{R}^p \times \mathbb{R}^d$ with joint distribution $P^{(X,Y)}$ is generated from

$$Y = b^*X + \varepsilon, \quad (1)$$

with regression coefficient $b^* \in \mathbb{R}^{d \times p}$, $\mathbb{E}X = 0$, and random residue $\varepsilon \in \mathbb{R}^d$ independent with X .

Objective: given $\{(X_i, Y_i)\}_{i=1}^n \stackrel{\text{iid}}{\sim} P^{(X,Y)}$, consider

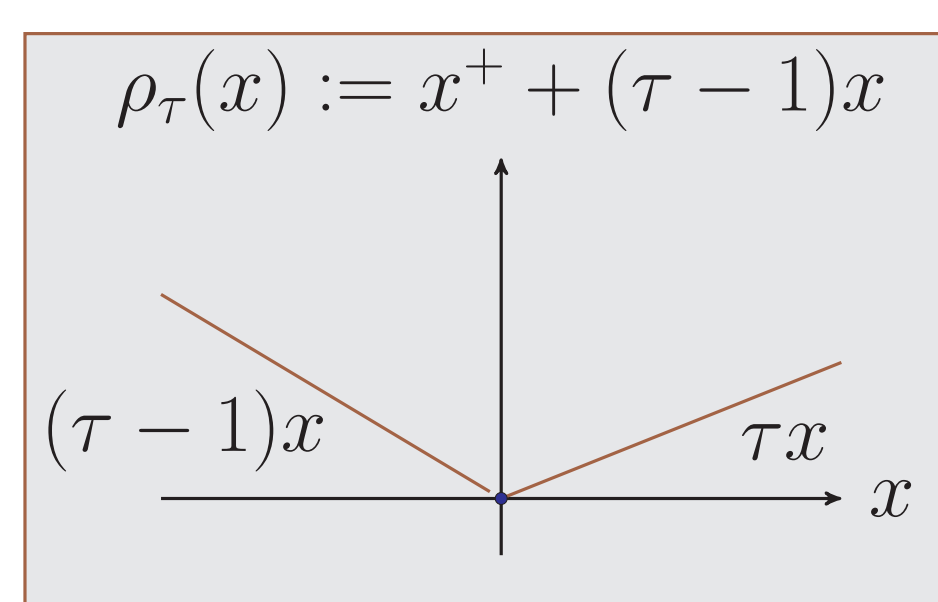
- heavy-tailed residue: $\varepsilon \sim P^\varepsilon$ allows ℓ -th moment (for $\ell > 2$);
- multiple-output response: $d \geq 2$.

We aim to estimate b^* .

When $d = 1$

Quantile regression: For a fixed $\tau \in (0, 1)$, consider

$$(\hat{b}, \hat{q}_\tau) = \arg \min_{\substack{b \in \mathbb{R}^{1 \times p} \\ q_\tau \in \mathbb{R}}} \sum_{i=1}^n \rho_\tau(Y_i - bX_i - q_\tau)$$



- Relative efficiency can be arbitrary samll!

Composite quantile regression (CQR): Let $\tau_k = k/(K+1)$, consider aggregated algorithm

$$(\hat{q}_1, \dots, \hat{q}_K, \tilde{b}) = \arg \min_{\substack{q_1, \dots, q_K \in \mathbb{R} \\ b \in \mathbb{R}^{1 \times p}}} \sum_{i=1}^n \sum_{k=1}^K \rho_{\tau_k}(Y_i - bX_i - q_k), \quad (2)$$

When $d \geq 2$: Check function ρ_τ is ill-defined!

Existing methods

- Projection method
 - Spatial quantile
 - The Monge-Kantorovich quantile
- Only capture convex support!

None of them is for robust coefficient estimation!

An optimal transport formulation

CQR in population formula Assume $q_1 \leq \dots \leq q_K$ in (2), then the corresponding population formula is

$$(b^*, q_\varepsilon^*) \in \arg \min_{\substack{b \in \mathbb{R}^{1 \times p} \\ q \in \mathcal{M}}} \mathbb{E} \int_0^1 \rho_\tau(Y - bX - q(\tau)) d\tau, \quad (3)$$

where \mathcal{M} denote the set of all increasing functions on \mathbb{R} . Let $U \sim \text{Unif}[0, 1]$, we have the following observation:

Start with CQR

$$\begin{aligned} & \inf_{q \in \mathcal{M}} \mathbb{E} \left\{ \int_0^1 \rho_\tau(Y - bX - q(\tau)) d\tau \right\} + \frac{1}{2} \mathbb{E}Y \\ &= \inf_{q \in \mathcal{M}} \left\{ \mathbb{E} \int_0^1 (Y - q(\tau) - bX)^+ d\tau + \int_0^1 (1 - \tau) q(\tau) d\tau \right\} \\ &= \inf_{q \in \mathcal{M}} \left\{ \mathbb{E} \max_{t \in [0,1]} \int_0^t (Y - q(\tau) - bX) d\tau + \mathbb{E} \phi(U) \right\} \\ &= \inf_{\phi \in \mathcal{C}} \left\{ \mathbb{E} \max_{t \in [0,1]} (t(Y - bX) - \phi(t)) + \mathbb{E} \phi(U) \right\} \end{aligned}$$

Conclusion: (3) $\Leftrightarrow b^* \in \arg \min_{b \in \mathbb{R}^{1 \times p}} \langle Y - bX, U \rangle_{\mathcal{W}_2}$.

- The OT formula can be extended to the case of $d > 2$ immediately by taking $U \sim \text{Unif}[0, 1]^d$!
- For any $P^\varepsilon, P^U \in \mathcal{P}_2(\mathbb{R}^d) \cap \mathcal{P}_{ac}(\mathbb{R}^d)$ and P^X is not a point mass, b^* is the unique minimiser, i.e.

$$b^* = \arg \min_{b \in \mathbb{R}^{d \times p}} \mathcal{L}(b), \quad \text{where } \mathcal{L}(b) := \langle Y - bX, U \rangle_{\mathcal{W}_2}$$

MCQR estimator Given $\{(X_i, Y_i)\}_{i=1}^n$ follows (1), and $\{U_i\}_{i=1}^m \stackrel{\text{iid}}{\sim} P^U$, the MCQR estimator is defined as

$$\hat{b} \in \arg \min_{b \in \mathbb{R}^{d \times p}} \mathcal{L}_{n,m}(b), \quad \text{where } \mathcal{L}_{n,m}(b) := \langle P_n^{Y-bX}, P_m^U \rangle_{\mathcal{W}_2}. \quad (4)$$

Theoretical guarantee

Let $P^U \sim \mathcal{N}(0, I_d)$ for theoretical convenience, and assume $P^\varepsilon \in \mathcal{P}_{ac}(\mathbb{R}^d)$ and P^X is an elliptical distribution.

Case 1: ε with finite ℓ -th moment ($\ell > 2$) Suppose $P^X, P^\varepsilon \in \mathcal{P}_\ell(\mathbb{R}^d)$, then with probability at least $1 - 4(\log n)^{-1}$, the MCQR estimator (4) satisfies

$$\|\hat{b} - b^*\|_{\Sigma}^2 \wedge 1 \leq C_{d,p} \left(n^{-\frac{1}{4}} + n^{-\frac{1}{d \vee p}} + n^{-\frac{\ell-2}{2\ell}} \right) \log m.$$

Case 2: ε with sub-Weibull tail:

- For $\sigma_1, \sigma_2 > 0$ and $\alpha, \beta \in (0, 2]$, $P^{\Sigma^{-1/2}X}$ is (σ_1, α) -sub-Weibull and P^ε is (σ_2, β) -sub-Weibull, i.e.

$$\mathbb{E} \exp\left\{ \frac{\|\Sigma^{-1/2}X\|/\sigma_1}{2} \right\} \leq 2 \quad \text{and} \quad \mathbb{E} \exp\left\{ \frac{\|\varepsilon\|/\sigma_2}{2} \right\} \leq 2$$

- For some $\gamma_1, \gamma_2 > 0$, the density of ε , write as f_ε satisfies $f_\varepsilon(e) \geq \gamma_1 \exp(-\gamma_2 \|e\|^2)$, for $\|e\| \geq 1$.

Then with probability at least $1 - 33(\log n)^{-1}$, we have

$$\|b^* - \hat{b}\|_{\Sigma}^2 \leq M_d \left((p/n)^{1/2} + n^{-2/d} \right) (\log m)^{\frac{8}{2\wedge\alpha\wedge\beta}}.$$

Proof Sketch

Case 1: Consider the basic inequality

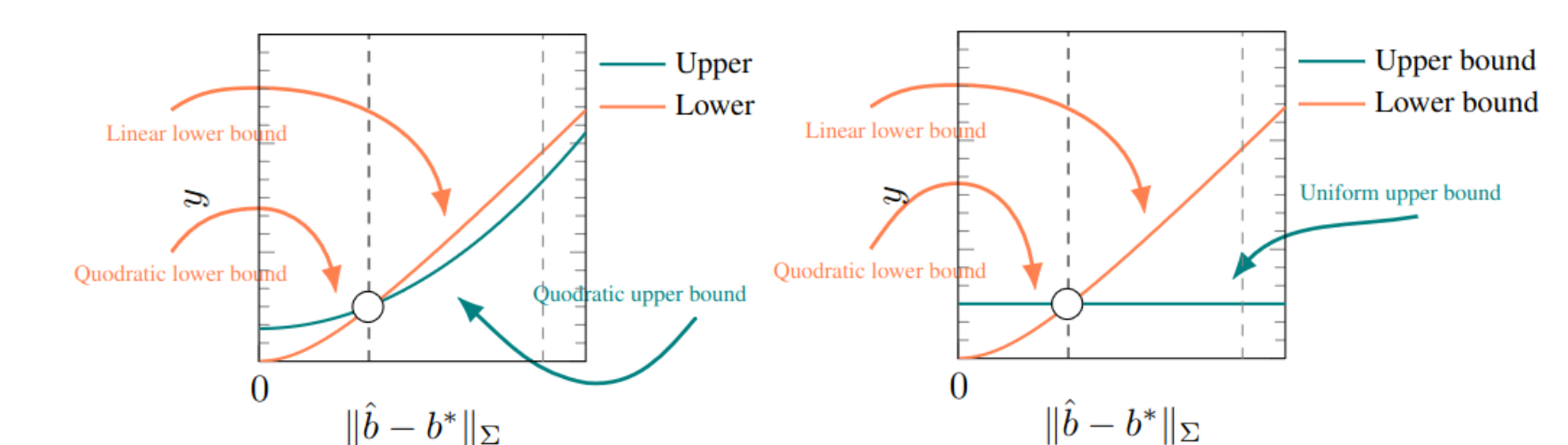
$$\mathcal{L}(\hat{b}) - \mathcal{L}(b^*) \leq \mathcal{L}(\hat{b}) - \mathcal{L}_{n,m}(\hat{b}) + \mathcal{L}_{n,m}(b^*) - \mathcal{L}(b^*).$$

- LHS Lower bound:

Lemma . Let $Z \perp \varepsilon$ random vectors in \mathbb{R}^d and $U \sim \mathcal{N}(0, I_d)$. If P^ε and P^Z are atomless with finite-second moments, then

$$\langle Z + \varepsilon, U \rangle_{\mathcal{W}_2}^2 \geq \langle Z, U \rangle_{\mathcal{W}_2}^2 + \langle \varepsilon, U \rangle_{\mathcal{W}_2}^2.$$

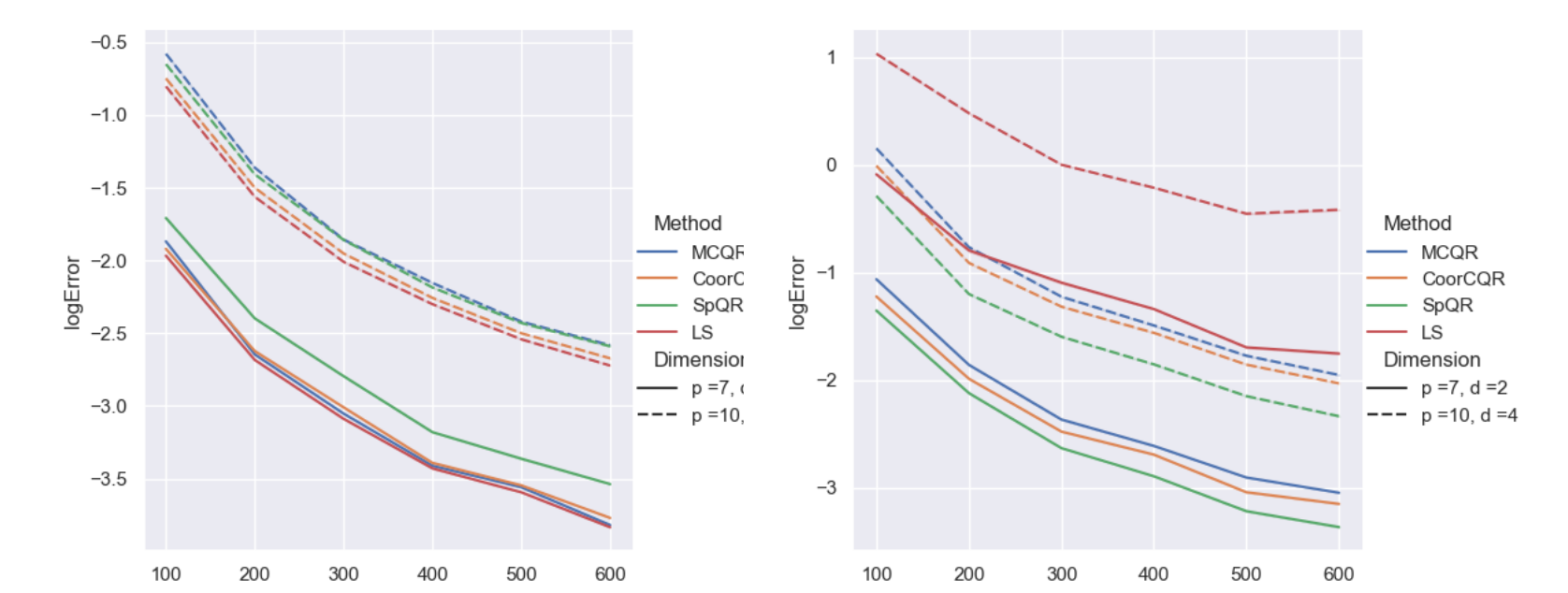
- RHS Upper bound



Case 2: Upper bound involves the following uniform error bound for empirical 2-Wasserstein distance:

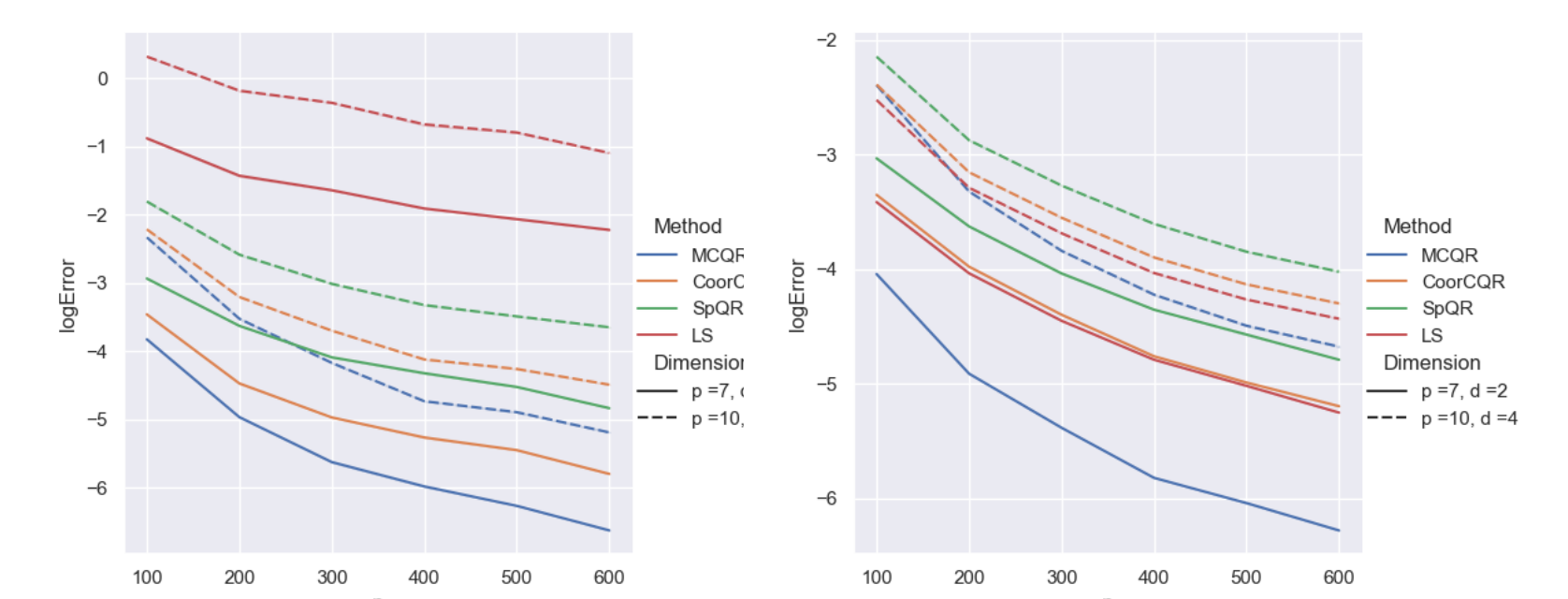
$$\sup_{b \in \mathcal{B}} \left| \mathcal{W}_2^2(P^{Y-bX}, P^U) - \mathcal{W}_2^2(P_n^{Y-bX}, P_m^U) \right|.$$

Experiments



(a) Gaussian noise

(b) Multivariate t_2



(c) Pareto copula

(d) Banana-shaped